Deleterious clinical effects of transfusion-associated immunomodulation: fact or fiction?

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Allogeneic blood transfusion results in the infusion into the recipient of large amounts of foreign antigens in both soluble and cell-associated forms. The persistence of these antigens in the circulation of the recipient may create conditions that allow the development of immune down-regulation. Evidence from a variety of sources indicates that allogeneic blood transfusion enhances the survival of renal allografts\(^1\) and may increase the recurrence rate of resected malignancies\(^2\) and the incidence of postoperative bacterial infections\(^3,7\), as well as reduce the recurrence rate of Crohn disease\(^8\) and/or activate infections with cytomegalovirus\(^9\) or human immunodeficiency virus.\(^10\) This clinical syndrome, the mechanisms of which remain to be defined, has been referred to in the transfusion medicine literature as allogeneic blood transfusion–associated immunomodulation (TRIM).\(^11\)

Clinical evidence for the existence of TRIM has been available since 1973. In their seminal study, Opelz et al\(^1\) provided evidence, counterintuitive at the time, that recipients of allogeneic blood transfusion had improved renal allograft survival. Subsequent clinical studies and studies in experimental animals corroborated the results of Opelz et al,\(^1\) and allogeneic blood transfusions were used deliberately in the early 1980s to prevent rejection of renal allografts.\(^12\) The beneficial effect of TRIM was obscured following the introduction of immunosuppression with cyclosporine, but it was recently reported to persist even in renal allograft recipients receiving modern immunosuppressive therapy.\(^13\)

On the basis of the immunomodulatory effect of allogeneic blood transfusion in renal allograft recipients, Gant\(^14\) raised the question in 1981 whether the TRIM effect might also be associated with an increased risk of cancer recurrence in patients undergoing resection of a malignancy. Gant’s hypothesis was based on the premise that, if allogeneic blood transfusion down-regulated the host’s immune surveillance mechanism that targets malignant cells, the receipt of allogeneic blood transfusion could enhance tumor growth. A subsequent hypothesis proposed that, if allogeneic blood transfusion causes immunosuppression, then recipients of perioperative allogeneic blood transfusion could be at increased risk for postoperative bacterial infection.

Since 1981, more than 150 clinical studies have examined the association of perioperative allogeneic blood transfusion with cancer recurrence and/or postoperative bacterial infection. Most of these are observational cohort studies comparing patients who had or did not have transfusion.\(^15-20\) In addition, 7 randomized controlled trials (RCTs) have compared the risk of cancer recurrence\(^2,21,22\) and/or postoperative infection\(^3,7,21-23\) between a treatment arm receiving standard\(^2\) or buffy-coat–reduced red blood cells (RBCs)\(^2,4,6,21,22\) or whole blood\(^3\) and a control arm receiving autologous or white blood cell (WBC)-reduced RBCs or whole blood.\(^24,25\) These studies are based on the assumption that the transfusion of autologous\(^2,21,23\) or WBC-reduced\(^5,7,22\) RBCs, or whole blood, is immunologically neutral. Both the earlier observational cohort studies and the recent RCTs have produced contradictory findings, and—because of the discrepancies among the published studies—the long-standing hypothesis of the potentially deleterious immunomodulatory effect of perioperative allogeneic blood transfusion remains unresolved.\(^11,26,27\)

The mechanism(s) of the TRIM effect(s) also remains elusive, and it is possible that a large number of biologic mechanisms may underlie the effect(s).\(^28-31\) The infusion of foreign antigen in either a soluble\(^31-36\) or cell-associated\(^37-43\) form has been shown to induce immune suppression, anergy, and clonal deletion in studies in experimental animals. However, most studies evaluating proposed mechanisms have been done in rodents, and these findings may not be applicable to the human immune system.\(^44\) Caution should be exercised, therefore, when findings from experimental animals are extrapolated to humans. Moreover, different biologic mechanisms may be involved in each reported clinical manifestation of TRIM.\(^1,10\) and the clinical evidence supporting each of the aforementioned hypotheses\(^1-10\) should be examined on its own merits.

The specific constituent(s) of allogeneic blood that mediates the TRIM effect\(^1,10\) remains unclear. Allogeneic plasma,\(^31-36\) allogeneic WBCs,\(^30,37-43\) and substances that accumulate in blood components during storage\(^39\) have been implicated in the pathogenesis of TRIM. However, both the animal and human data suggest that the TRIM effects are most likely mediated by transfused allogeneic WBCs.\(^5,46\) Lee et al\(^47\) reported persistence of donor WBCs in humans for up to 1.5 years after an allogeneic blood transfusion. In murine and rabbit experimental models, Blajchman et al\(^37,38\) and Bordin et al\(^39\) demonstrated a tumor growth–promoting effect of allogeneic blood transfusion that appeared to be associated with transfusion of allogeneic WBCs. The findings of these experiments support the hypothesis that allogeneic WBCs actively induce immunosuppression in transfusion recipients. In another relevant study, Kao\(^40\) induced immune suppression in mice receiving donor WBCs free of plasma and platelets. These and several other investigators\(^41,42\) also attributed an induction of T\(_H\)2 cells in transfusion recipients to the allogeneic donor WBCs, showing that T\(_H\)2 cells can produce immunosuppression in transfusion recipients by down-regulating the activity of T\(_H\)1 cells. Minchell et al\(^43\) implicated the antigen-presenting cells of the allogeneic donor in the induction of a state of anergy in the recipient and proposed that, during refrigeration, antigen-presenting cells lose their ability to deliver costimulation. These investigators hypothesized that after an allogeneic transfusion, the recipient’s T cells are stimulated by

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allogeneic donor WBCs in the absence of costimulation and that this interaction induces a state of anergy in the recipient’s T cells.

More than 100 observational cohort studies comparing patients with and without transfusion, as well as 2 RCTs comparing recipients of allogeneic and autologous RBCs, did not associate the TRIM effect with any particular blood constituent. Five recent RCTs were designed on the basis of the assumption that the allogeneic transfusion effect responsible for cancer recurrence and/or postoperative infection is mediated specifically by allogeneic WBCs. These studies produced contradictory findings, and the issue as to whether any deleterious TRIM effects are lessened or abrogated by WBC reduction of the transfused allogeneic cellular blood components remains unresolved.11,26,27,48 Jensen et al attributed an 86% reduction in the incidence of postoperative infection to the use of WBC-reduced allogeneic whole blood or RBCs,5 with the use of poststorage WBC reduction by filtration. In contrast, Houbiers et al found no difference in the incidence of postoperative infection between the recipients of WBC-reduced and buffy-coat–reduced RBCs with the use of poststorage WBC reduction by filtration. In the RCT by van de Watering et al,6 the use of WBC-reduced RBCs decreased the incidence of postoperative infection by 30% compared with that seen in patients having transfusion with buffy-coat–reduced RBCs. Moreover, there was no difference in the incidence of postoperative infection between the recipients of RBCs that were WBC reduced by filtration before or after storage.

Recently, the United Kingdom, Ireland, and Portugal implemented universal WBC reduction of all transfused cellular blood components, based on the hypothesis that this practice would prevent the theoretical risk of transmission by transfusion of the agent of variant Creutzfeldt-Jakob disease (vCJD). France and Canada also implemented universal WBC reduction, primarily to enhance overall transfusion safety. Following these developments, public debate began regarding the appropriateness of introducing universal WBC reduction in the United States and elsewhere. The risk of transmission of vCJD by transfusion is considered only a theoretical possibility because an epidemic of bovine spongiform encephalopathy and cases of vCJD have occurred almost exclusively in the United Kingdom. Therefore, if a decision were made to implement universal WBC reduction in the United States and elsewhere, such a policy would probably not be introduced with an intent to prevent the transmission of vCJD, but to reap other potential benefits of WBC reduction. Most of the published evidence regarding such potential benefits pertains to the prevention of the deleterious TRIM effects, and the public debate regarding the implementation of universal WBC reduction is likely to focus mainly (although not exclusively) on an examination of the efficacy of WBC reduction in preventing these potential adverse effects of allogeneic blood transfusion. Recently, the described relation between WBC-containing allogeneic blood transfusion and increased postoperative mortality from causes other than postoperative infection, and the report by Hébert et al that a restrictive strategy of RBC transfusion may be superior to a liberal transfusion strategy when mortality is evaluated as an outcome in critically ill patients, may affect the debate over whether to implement universal WBC reduction.

Thirty-one physicians from academic blood banks in the United States recently voiced their strong opposition to the intent of the Food and Drug Administration (FDA) to mandate the implementation of universal WBC reduction in the United States. In a letter to the editor of Transfusion, they stated that “It is our view that published reports fail to document a substantial health benefit that would serve to justify WBC reduction of cellular blood components transfused to all patients. Accordingly, we feel that the currently available evidence regarding the deleterious effects of allogeneic blood transfusion is not sufficiently compelling to warrant universal WBC reduction for the prevention of these effects.” This present review examines the available evidence from: (1) the observational cohort studies that investigated the hypothesis that allogeneic blood transfusion provokes cancer recurrence and/or postoperative infection; (2) the findings of the RCTs that examined the possible deleterious TRIM effects associated with allogeneic blood transfusion in general or with transfusion of allogeneic WBCs in particular; and (3) the results of the recent RCTs that reported an association between allogeneic blood transfusion and increased mortality. To help readers process and understand the contradictory information, we include a section on the rationale, design, and analysis of the clinical studies that investigated these possible deleterious TRIM effects. Also, after reviewing the available evidence regarding the association of allogeneic transfusion with cancer recurrence or postoperative infection, we discuss the various arguments for and against a potential decision to implement universal WBC reduction of all transfused cellular blood components to prevent the deleterious effects of allogeneic transfusion-associated immunomodulation.

**Design of clinical studies**

Evidence in support of deleterious TRIM effects has been obtained from observational cohort studies, RCTs, meta-analyses of the observational cohort studies, and meta-analyses of the RCTs. The observational studies were either retrospective or prospective, and compared the risk of an adverse clinical outcome between cohorts of patients who did or did not receive transfusion and who differed with respect to baseline determinants of the need for transfusion and the risk of cancer recurrence or postoperative infection. These studies relied on multivariate statistical analyses to adjust for the effects of possible confounding factors. The RCTs were prospective clinical experiments comparing the risk of an adverse outcome between patients randomly allocated to receive different blood products, thus relying on randomization (ie, chance) to distribute all possible confounding factors equally between the treatment and control arms. The RCTs presented univariate (ie, unadjusted) analyses calculating the odds ratio (OR) of an adverse outcome in a treatment arm compared with that in a control arm, as shown in Table 1. Such analyses are valid if the counts in each of the 4 cells of the 2 × 2 contingency table (Table 1) are free of the effects of confounding factors—a condition usually satisfied when large numbers of patients are randomized. Meta-analyses of RCTs integrate the ORs from individual RCTs into a summary OR, which is also free of the effects of confounders and represents a measure of the “average” TRIM effect across the combined RCTs. Meta-analyses of observational studies extract 2 × 2 contingency table counts from each study, use these 4 counts to calculate an OR for each corresponding original report, and then also integrate the ORs from the individual studies into a summary OR. Unlike the summary ORs calculated by combining the results of RCTs, the summary ORs reported from meta-analyses of observational studies incorporate the effects of all known and unknown confounders of the association of transfusion with cancer recurrence and/or postoperative infection. Therefore, the summary ORs calculated by meta-analyses of the observational studies of the adverse TRIM effects should not be interpreted as a measure of the “average” TRIM effect across the combined studies. In only one of these meta-analyses was meta-regression used to adjust for the effects
of 3 confounders on which information had been reported by most of the published observational studies.

When the results of studies retrieved for a meta-analysis are discrepant, or if the variation among the reported findings is greater than can be attributed to chance (a situation called “heterogeneity of effects”), a meta-analysis can offer insight into the reasons for the disagreements among the available studies. Provided that the results from the available studies are concordant, meta-analysis can be used to derive, through the application of a number of quantitative techniques, a measure of the effect of allogeneic blood transfusion based on a combination of the available data. This “average” TRIM effect is more precise (and more likely to attain statistical significance) than the TRIM effects reported from individual studies because it is based on a much larger sample (ie, the study population accrued when the study populations from all available reports are combined). Before the results of studies are integrated, however, the degree of agreement among reports must be assessed both conceptually and statistically. The Q test statistic quantifies the probability that the variation in reported results might have arisen by chance. Most analysts hold that there is sufficient agreement (or “homogeneity”) among studies to permit the undertaking of a meta-analytic synthesis of the results if $P > .05$ for the Q test statistic (ie, if there is a greater than 5% probability that the variation in reported results might have arisen by chance). In this review, we use meta-analysis both to investigate reasons for disagreements among studies (if $P < .05$ for the Q test statistic) and to integrate the results of the available studies for the calculation of an “average” TRIM effect across combined investigations (if $P > .05$ for the Q test statistic). Where meta-analysis is used in this review, results of homogeneous studies are integrated using the random-effects method of DerSimonian and Laird.

Patients allocated to the treatment arm of the reviewed RCTs were exposed to all the constituents of allogeneic blood and were at risk for developing adverse TRIM effects. Control-arm patients were presumed to be at reduced risk because they were not exposed to allogeneic WBCs or allogeneic WBCs and allogeneic plasma. 

### Table 1. Calculation of the odds ratio for studies investigating the relationship between allogeneic blood transfusion and cancer recurrence or postoperative infection

<table>
<thead>
<tr>
<th>Group</th>
<th>Adverse outcome: cancer recurrence or postoperative infection</th>
<th>Occurring</th>
<th>Not occurring</th>
<th>Totals</th>
</tr>
</thead>
<tbody>
<tr>
<td>Exposed to allogeneic transfusion*</td>
<td>a</td>
<td>b</td>
<td>$a + b$</td>
<td></td>
</tr>
<tr>
<td>Not exposed to allogeneic transfusion†</td>
<td>c</td>
<td>d</td>
<td>$c + d$</td>
<td></td>
</tr>
<tr>
<td>Totals</td>
<td>a + c</td>
<td>b + d</td>
<td>$a + b + c + d$</td>
<td></td>
</tr>
</tbody>
</table>

Odds of an adverse outcome in those exposed to allogeneic transfusion: $a / (a + b)$
Odds of an adverse outcome in those not exposed to allogeneic transfusion: $c / (c + d)$

$$
\text{Odds ratio} = \frac{\text{Odds of an adverse outcome in those exposed to allogeneic transfusion}}{\text{Odds of an adverse outcome in those not exposed to allogeneic transfusion}}
$$

### Table 2. Effect of buffy-coat reduction in decreasing allogeneic blood transfusion–induced tumor growth in New Zealand White rabbits with established tumors

<table>
<thead>
<tr>
<th>Red cell transfusion received</th>
<th>Sample size</th>
<th>Number of pulmonary tumor nodules (median ± SE)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Allogeneic, unmodified (stored)</td>
<td>20</td>
<td>74.0 ± 8.7</td>
</tr>
<tr>
<td>Allogeneic, buffy-coat reduced before storage†</td>
<td>19</td>
<td>34.0 ± 7.0</td>
</tr>
<tr>
<td>Allogeneic, WBC reduced before storage‡</td>
<td>20</td>
<td>23.5 ± 1.8</td>
</tr>
<tr>
<td>Control group (no transfusion)</td>
<td>16</td>
<td>21.5 ± 2.5</td>
</tr>
</tbody>
</table>

WBC indicates white blood cell.

*Modified from Blajchman.
†Mean decrease in WBC dose of 76% (similar to the buffy-coat reduction method used in the RCTs reported from western Europe).
‡Mean decrease in WBC dose of 99.7%.
blood\textsuperscript{3} to the control arm transfused RBCs prepared by leukocyte filtration. Prestorage-filtered RBCs were prepared by passing, within 24\textsuperscript{h} or 48\textsuperscript{h} hours of collection, a buffy-coat–reduced RBC unit through a leukocyte-reduction filter. Poststorage-filtered RBCs were prepared by passing a unit of buffy-coat–reduced RBCs\textsuperscript{5,9} or whole blood,\textsuperscript{3} stored for 6 or more days, through a leukocyte-reduction filter. The average WBC content of transfused WBC-reduced RBCs was less than 5 \times 10^\textsuperscript{3} per unit, and was often as low as 1 \times 10^\textsuperscript{3} per unit. The timing of the filtration procedure is important because biologic response modifiers released from WBC granules in a time-dependent manner during storage may mediate the adverse TRIM effects,\textsuperscript{39,48,56-60} and—if such TRIM effects really exist—only WBC-reduced RBCs filtered before storage can be expected to abrogate the adverse effects. Two RCTs administered autologous RBCs\textsuperscript{4} or whole blood\textsuperscript{31} to the control arm. Autologous blood can be expected to abrogate the TRIM effects mediated by allogeneic plasma,\textsuperscript{31-36} or by intact, immunologically competent allogeneic WBCs, but not the TRIM effects mediated by biologic response modifiers released from WBC granules during storage.\textsuperscript{59}

**Observational studies**

Observational studies have compared the incidence of cancer recurrence, death due to cancer recurrence, and/or overall mortality\textsuperscript{15-18} between patients undergoing cancer resection who did or did not receive transfusion; or the incidence of postoperative bacterial infection with or without transfusion in patients undergoing gastrointestinal surgery, orthopedic operations, cardiac surgery, or various other procedures. These studies tended to indicate that patients having transfusion (compared with those not having transfusion) had a higher incidence of cancer recurrence or death due to cancer recurrence as well as a shorter overall survival after a cancer resection operation; and almost always had a higher incidence of postoperative bacterial infection.\textsuperscript{16} These studies also indicated that patients having transfusion generally differed from those without transfusion in several prognostic factors, including clinical stage of the malignancy; size, histologic grade, and type of tumor; age; preoperative hemoglobin; duration and extent of surgery; amount of perioperative blood loss; the frequency of chronic systemic illness, such as congestive heart disease, lung disease, liver disease, kidney failure, or diabetes mellitus; and the presence of risk factors for postoperative urinary tract infection (UTI), pneumonia, or wound infection.\textsuperscript{18,19,61}

These 2 sets of observations have led to different interpretations of the results of the studies that compared patients who did or did not receive transfusion. Some authors concluded that perioperative allogeneic blood transfusion had a direct deleterious effect on the recipient, causing an increase in the incidence of cancer recurrence and/or postoperative bacterial infection.\textsuperscript{16} Other investigators concluded that allogeneic blood transfusions were a surrogate marker for a variety of adverse prognostic factors and that the other variables that generated the need for perioperative transfusions determined the subsequent clinical outcome.\textsuperscript{19} The latter school of thought reasoned that the adverse prognostic factors that were associated with both the need for perioperative transfusion and the cancer recurrence and/or postoperative bacterial infection were confounders of the relation between transfusion and these outcomes, and engendered a spurious association between allogeneic transfusion and cancer recurrence and/or postoperative infection.\textsuperscript{19}

In many of the reported observational studies, the authors used multivariate regression analysis to adjust the effect of transfusion for the effects of confounding factors. Regression analysis can calculate an allogeneic blood transfusion effect that is independent of the effects of known and measurable confounders, provided that: (1) the investigators measure all known potentially confounding factors; (2) the variables that are associated with both transfusion and cancer recurrence and/or postoperative infection in a particular set of data are identified as true confounders; and (3) all true confounders are included in the final regression model on which the conclusions are based.\textsuperscript{19} In most published observational studies, however, important potential confounding factors were not considered by the investigators\textsuperscript{61}; and the multivariate regression model on which the conclusions were based was built by a stepwise (as opposed to forced-entry) method, which did not ensure the inclusion of several measured confounders in the final model.\textsuperscript{19} For example, although postoperative UTI was the most frequent type of postoperative infection in most observational studies that reported an association between transfusion and postoperative infection, the number of patient days with an indwelling urinary catheter—the predominant risk factor for nosocomial UTI in the United States—was not considered as a potential confounder and was not measured in most studies.\textsuperscript{61}

As a result, TRIM effects reported as “independent” by many teams of investigators were not truly free of the effects of known and measurable confounding factors, and the calculated TRIM effect decreased in magnitude (or became statistically insignificant) depending on which confounding factors were included in the final regression model.\textsuperscript{61} These caveats notwithstanding, allogeneic blood transfusion often emerged as the leading predictor of cancer recurrence and/or postoperative bacterial infection from multivariate analyses, and was reported to increase the risk of postoperative infection in patients having transfusion (compared with those not having transfusion) by up to 10-fold.\textsuperscript{16}

**Cancer recurrence**

Chung et al,\textsuperscript{62} Vamvakas,\textsuperscript{17} and Brand and Houbiers\textsuperscript{18} used meta-analysis to integrate the findings from the univariate (ie, unadjusted; Table 1) analyses of the observational studies of the association of perioperative allogeneic blood transfusion with cancer recurrence, death due to cancer recurrence, or overall mortality. The most recent meta-analysis\textsuperscript{18} also included studies not published in English and considered 7 cancer sites; 100 observational studies were identified as eligible for analysis. There was agreement among the 3 meta-analyses with regard to the magnitude and statistical significance of the calculated summary ORs of the risk of an adverse clinical outcome in patients having transfusion (compared with those not having transfusion). Figure 1 shows the unadjusted summary OR of an adverse outcome for 7 cancer sites. As shown in Figure 1, when the unadjusted results were integrated from 28 reports on colorectal cancer, 14 reports on head and neck cancer, 10 reports on breast cancer, 8 reports on gastric cancer, 8 reports on lung cancer, 6 reports on cervical cancer, and 6 reports on prostate cancer, the summary OR of an adverse clinical outcome with transfusion (versus without transfusion) was statistically significant (P \leq 0.05) for all sites except the cervix (P > 0.05). These unadjusted results contrast with the conclusions of most of the studies that presented a multivariate analysis. A statistically significant (P \leq 0.05) TRIM effect adjusted for the effects of confounding factors was reported from only 11 studies of colorectal cancer,\textsuperscript{17} 4 studies of head and neck cancer, 1 study of breast cancer, 2 studies of gastric cancer, 4 studies of lung cancer, and 2 studies of prostate cancer.\textsuperscript{18}

On the basis of such a statistical combination of the unadjusted findings of the retrieved observational studies, Chung et al\textsuperscript{62}
concluded that allogeneic blood transfusion given for a cancer resection operation had an adverse effect on subsequent prognosis; Vamvakas and Brand and Houbiers reached the opposite conclusion. In the case of colorectal cancer studies, there was marked variation among the findings of the available reports, and the probability that such variation might have arisen by chance was smaller than 1 per 1000 (P < .001 for the Q test statistic). As already discussed, this situation precludes the calculation of a summary estimate of the allogeneic transfusion effect across the available studies by the techniques of meta-analysis.59

Brand and Houbiers separately analyzed the studies that reported on colorectal cancer recurrence (27 studies), death due to colorectal cancer recurrence (15 studies), or overall mortality following colorectal cancer resection (30 studies). When the unadjusted findings from these reports were integrated, there was still unexplained heterogeneity (P < .001 for the Q test statistic) in all 3 analyses. Meta-regression was used to adjust the summary OR of cancer recurrence for the effects of tumor location and clinical (Dukes) stage of tumor; the summary OR of death due to cancer recurrence for the effect of tumor location; and the summary OR of death from any cause for the effects of tumor location, clinical (Dukes) stage of tumor, and mean patient age. After adjustment for the effects of these confounding factors, no association was found between perioperative transfusion and cancer recurrence (P = .55) or death due to cancer recurrence (P = .19), but the relationship between perioperative transfusion and overall mortality persisted (P < .001). Brand and Houbiers therefore reasoned that, because meta-regression could attribute the association of transfusion with cancer recurrence (or death due to cancer recurrence) to the effects of confounding factors, and because only a minority of the published observational studies had reported a significant adverse TRIM effect based on multivariate analyses, a link between perioperative transfusion and colorectal cancer recurrence could not be established. These investigators concluded that the observed effect of allogeneic transfusion on overall mortality might be due to an association of perioperative transfusion with causes of death other than cancer recurrence.18

It is also possible to ascribe the observed effect of allogeneic transfusion on mortality to a greater physiologic severity of illness in patients having transfusion than in those not having transfusion. Most observational studies presented data on the severity of disease in the included subjects; that is, they reported the clinical (Dukes) stage of tumor. In contrast, severity of illness incorporates both the severity of disease and the prevalence and severity of comorbidities, such as diabetes mellitus, congestive heart disease, lung disease, liver disease, or kidney failure. Such chronic systemic illnesses may have as much of an impact on overall mortality as the clinical (Dukes) stage of tumor, but the available observational studies did not present data on chronic systemic illness, making it impossible to adjust for the effect of illness severity in the meta-analysis. It is thus possible that the patients having perioperative transfusion survived for a shorter period after the cancer resection operation simply because they were sicker at the time of surgery than the patients who did not require transfusion.

The findings of the available observational studies on the association of perioperative transfusion with cancer recurrence may no longer be relevant to the management of contemporary cancer patients. Many of these studies (eg, all of the studies that reported on patients with gastric cancer) examined only overall survival; others did not report a multivariate analysis; still others included patients operated on before 1980 and over an extended period. For example, only 3 of the 14 studies of head and neck cancer retrieved for the meta-analysis of Brand and Houbiers fulfilled the selection criteria of reporting on cancer-specific complications rather than overall mortality, adjusting for the effects of confounding factors by multivariate analysis, and including patients operated on after 1980 and over a period of less than 5 years. Only 1 of these 3 studies reported a statistically significant adverse effect of allogeneic blood transfusion.18

Furthermore, the findings of the available observational studies of the association of perioperative transfusion with cancer recurrence may reflect the effect of publication bias, a consideration that may also apply to the discussion of the reported observational studies of the relationship between perioperative transfusion and postoperative infection. Studies reporting null results are less likely to be published than studies reporting statistically significant findings. Easterbrook et al documented a statistically significant 3.8-fold increase in the odds of publication for observational studies reporting significant findings, compared with studies with null results. Multivariate analysis showed that the better odds of publication could not be explained by the quality of the study design. On the contrary, there was a trend toward a greater number of significant results with poorer quality studies.

Rather than comparing patients with or without transfusion, Ness et al prospectively compared the recurrence rate of prostate cancer in 309 recipients of autologous or allogeneic blood transfusion, and observed no clinical benefit from the use of autologous blood. The hazard ratio of allogeneic transfusion in a univariate Cox proportional hazard model of time to tumor recurrence was 0.87 (P > .05), and Ness et al commented that a beneficial TRIM effect could not be excluded. It is noteworthy that allogeneic blood transfusion is not associated with an adverse prognosis in patients with cervical cancer, even when all available univariate (unadjusted) results are combined (Figure 1). Because human papilloma-virus (HPV) is implicated in the pathogenesis of cervical cancer, cytotoxic T cells directed against HPV antigens expressed on the surface of tumor cells could contribute to immunologic control of the growth of residual tumor cells following cancer resection.

Figure 1. Summary odds ratio of an adverse clinical outcome (ie, cancer recurrence, death due to cancer recurrence, or overall mortality) across published observational studies comparing patients having or not having transfusion. The univariate (unadjusted) results of the available studies are integrated separately for each cancer site. The data on colorectal and prostate cancer (representing 28 and 6 studies, respectively) are shown as reported by Vamvakas and Brand and Houbiers. The data on head and neck (14 studies), breast (10 studies), gastric (8 studies), lung (8 studies), and cervical cancer (6 studies) are shown as reported by Brand and Houbiers. The references to the included primary studies can be found in the reports of these 2 meta-analyses. Each summary OR is surrounded by its 95% CI. A summary OR of the null value (1) indicates that the risk of an adverse clinical outcome is the same, on average, with or without transfusion across the combined studies. When the 95% CI of the summary OR extends on both sides of the null value, the calculated summary TRIM effect is not statistically significant.
Accordingly, if allogeneic blood transfusion induced immune suppression, it could down-regulate these cytotoxic T-cell responses, promoting tumor cell growth. Transfusion would thus be expected to have a larger adverse effect in patients with cervical cancer than in patients with cancer from other sites (Figure 1) where the tumors are not virus induced and are probably subject to weaker immunologic control mechanisms. However, it has been observed that early in the development of cervical intraepithelial neoplasia, the expression of class I HLA molecules on malignant cells is specifically down-regulated. HPV-specific, HLA-restricted cytotoxic T cells (whose function may be suppressed by allogeneic blood transfusion) may not be effective against cervical cancer cells because of the reduction in class I HLA molecule expression on the tumor cells.

**Postoperative infection**

Except for a handful of studies, observational investigations of the association of allogeneic blood transfusion with postoperative bacterial infection almost uniformly have reported a relationship between perioperative transfusion and infection, which persisted after statistical adjustment for the effects of the confounding factors considered by the authors. Postoperative infection often develops, however, because of a higher physiologic severity of illness and a higher prevalence of risk factors for postoperative infection at specific sites in patients having transfusion compared with those not having transfusion. As already discussed in the context of cancer recurrence, in most published observational studies of transfusion and postoperative infection, the reported allogeneic transfusion effect was adjusted for the effect of severity of disease (ie, the severity of the principal diagnosis), but the severity of the principal diagnosis differs from the severity of a patient’s overall illness. In a study that compares the frequency of postoperative infection with or without transfusion in patients undergoing colorectal cancer resection, the presence and severity of comorbidities (eg, diabetes mellitus, congestive heart disease, lung disease, liver disease, or kidney failure) may be a more important determinant of postoperative infection than the severity of the principal diagnosis according to the clinical (Dukes) stage of tumor. Furthermore, the number of patient days with an indwelling urinary catheter may be the most important determinant of postoperative UTI; the number of days of endotracheal intubation and/or impaired consciousness may be a cardinal determinant of postoperative pneumonia; and so on.

Until recently, observational studies reporting an association between allogeneic blood transfusion and postoperative infection did not adjust for the effects of severity of illness and/or risk factors for postoperative infection at specific sites. Some teams of investigators secured partial control for the effects of these variables by excluding UTIs from the definition of postoperative infection; by limiting the outcome variable to postoperative wound infection; or by adjusting for factors of serum albumin, insertion of a urinary catheter, or presence of chronic systemic illness. However, adjustment for the effects of all these factors in combination has rarely been presented in the literature.

In a study of 492 patients undergoing colorectal cancer resection, Vamvakas et al calculated the probability of infection in association with perioperative blood transfusion with and without adjustment for the effects of chronic systemic illness, number of days with an indwelling urinary catheter, endotracheal intubation, and impaired consciousness. In an analysis that adjusted only for the effects of 18 confounding variables considered by previous authors and that adjusted for insertion of a urinary catheter (as opposed to number of patient days with an indwelling urinary catheter), these investigators detected a highly significant (P < .001) transfusion effect on the risk of postoperative infection at any site. However, when adjustment was also made for the effects of the aforementioned variables, the association between transfusion and postoperative infection at any site disappeared (P = .407); the only significant predictors of postoperative infection were the number of patient days with an indwelling urinary catheter, the presence of chronic systemic illness, the number of days of impaired consciousness, and the duration of anesthesia. Except for the duration of anesthesia, the 17 other confounding variables considered by previous authors proved to be insignificant predictors of postoperative infection in this analysis.

Observational studies of the association of allogeneic blood transfusion with postoperative infection in patients having orthopedic surgery secured partial adjustment for the effect of greater illness severity in the transfusion group by comparing subjects who received autologous or allogeneic blood transfusion. At least to some extent, patients who made preoperative autologous blood donations did so because they were in better health than subjects who did not make autologous donations. Also, autologous blood may have been transfused more liberally than allogeneic blood, after less surgical blood loss, and it is difficult to attribute any increase in the infection rate seen in the allogeneic (compared with the autologous) transfusion group to the allogeneic transfusion per se if adjustment is not made for the effects of chronic systemic illness and risk factors for postoperative infection at specific sites. Duffy and Neal conducted a meta-analysis of the univariate (ie, unadjusted) results of 5 observational studies and 2 RCTs comparing the postoperative infection rates of patients having transfusion with similar volumes of autologous or allogeneic blood. Patients receiving no transfusion or a mixture of autologous and allogeneic RBCs in these studies were excluded from the meta-analysis. Because it is uncommon for more than 3 predonated units of autologous blood to be available, subjects receiving 4 or more units of allogeneic RBCs were also excluded. The studies of Busch et al and Triulzi et al were excluded because of insufficient published information. The summary OR across the 7 studies was 2.4 (95% confidence interval [CI], 1.6-3.6; P < .0001). Figure 2 updates the meta-analysis of Duffy and Neal by including the study of Innerhofer et al. The 8 available studies were homogeneous (P = .50 for the Q test statistic), and the summary OR of postoperative infection in the allogeneic (compared with the autologous) transfusion group was 2.1 (95% CI, 1.3-3.4; P < .001).

Recently, Carson et al conducted a retrospective cohort study of 9598 consecutive patients with hip fracture who underwent surgical repair between 1983 and 1993 at 20 hospitals across the United States. The primary outcome variable was serious bacterial infection, defined as bacteremia, pneumonia, deep wound infection, or septic arthritis/osteomyelitis. Information was collected on numerous variables, including comorbid conditions such as the determinants of the Charlson Comorbidity Index, but the method used for building the statistical models was not described. The adjusted relative risk of serious postoperative infection with transfusion (versus without transfusion) was 1.43 (95% CI, 1.16-1.78; P = .001). Chang et al analyzed a database of 1349 patients undergoing elective colorectal surgery for any disease of the colon or rectum at 11 university hospitals across Canada. To better adjust for the effects of factors confounding the association of transfusion with postoperative infection, these investigators limited the outcome variable to postoperative wound infection. Ten prognostic variables were found to be associated with both transfusion and postoperative wound infection, and the final regression model...
adjusted for 4 of these identified confounders. Allogeneic blood transfusion was reported to be a significant independent predictor of postoperative wound infection (OR = 1.18; 95% CI, 1.05-1.33; P = .007). Vamvakas and Carver reported a retrospective cohort study of 416 consecutive patients admitted to one hospital for coronary artery bypass graft (CABG) operations. The outcome variable was limited to postoperative wound infection or pneumonia, and adjustment was made for the effects of chronic systemic illness and specific risk factors for wound infection or pneumonia. Statistical models were built by the forced-entry method, and the adjusted risk of postoperative wound infection or pneumonia increased by 6% per unit of allogeneic RBCs and/or platelets transfused (P = .0284), or by 43% for a patient receiving the mean transfusion dose of 7.2 U of RBCs and/or platelets.

Randomized controlled trials

Cancer recurrence

The 3 RCTs that compared the incidence of cancer recurrence between recipients of buffy-coat-reduced allogeneic RBCs and recipients of autologous whole blood, or WBC-reduced, buffy-coat-reduced allogeneic RBCs were medically and statistically homogeneous. All 3 studies enrolled patients undergoing colorectal cancer resection. The proportion of patients having transfusion varied from 58% to 64% among the studies, and the proportion of patients developing recurrent cancer varied from 23% to 25.5%. There was also agreement among the findings of these RCTs (Figure 3), as the noted variation in reported results was sufficiently modest to be attributed to chance (P > .10 for the Q test statistic). Accordingly, the findings of the studies were combined in 2 meta-analyses, and the summary OR of cancer recurrence in the allogeneic transfusion (compared with the control) group across the 3 studies was 1.04 (95% CI, 0.81-1.35; P > .05) (Figure 3). The summary OR of death due to cancer recurrence was 0.98 (95% CI, 0.76-1.26; P > .05).

Busch et al reported on 423 of 510 randomized patients; Houbiers et al reported on 697 of 1021 randomized subjects, and Heiss et al reported on 100 of 120 randomized patients. In addition to the patients who did not have transfusion (42%, 36%, and 40% of the study samples, respectively), 28%, 10%, and 33%, respectively, of the subjects randomized to receive autologous or WBC-reduced allogeneic RBCs also received buffy-coat-reduced allogeneic RBCs. Many violations of the experimental protocol occurred in the studies of Busch et al and Heiss et al because of a design problem that needs to be discussed. Patients randomly allocated to the autologous transfusion arm in these studies donated 2 U of whole blood before surgery, and, if they needed transfusion of more than 2 U of RBCs perioperatively, they were given buffy-coat-reduced RBCs. In addition, because of the preoperative autologous blood donations, these patients presented to the operating room with a lower hematocrit than patients from the allogeneic transfusion arm. Therefore, autologous transfusion recipients probably received transfusions sooner, after less surgical blood loss, than did patients allocated to the allogeneic transfusion arm. It is thus possible that patients from the autologous arm having transfusion with a particular number of RBC units may have had less invasive surgery than patients from the allogeneic transfusion arm given the same number of RBCs; this design problem may have led to overestimation of any adverse effect of allogeneic blood transfusion, as discussed by Heiss et al.

It is impossible, for ethical reasons, to perform RCTs in which patients are randomly allocated not to receive blood transfusion or to always receive transfusion; only patients prospectively randomized to receive different blood products—should the need for transfusion arise perioperatively—can be compared in RCTs. Ideally, the blood components administered to patients enrolled in RCTs should differ in only one factor, which reflects a biologic mechanism presumed to underlie the immunosuppressive effect of allogeneic blood transfusion. If the TRIM effect is assumed to be mediated by allogeneic WBCs, recipients of standard or WBC-reduced RBCs should be compared. However, only one RCT investigating the association between allogeneic transfusion and...
cancer recurrence has transfused WBC-reduced RBCs to patients from the control arm.

Because of the high frequency of patients who did not have transfusion and the protocol violations in the 3 RCTs investigating the association of allogeneic transfusion with cancer recurrence, a deleterious TRIM effect on colorectal cancer recurrence could be evaluated in an effective sample of only 696 patients across the 3 studies.18 Therefore, the meta-analysis of the 3 RCTs did not have adequate statistical power to rule out the possibility of an adverse TRIM effect smaller than a 33% increase in the risk of cancer recurrence among the recipients of buffy-coat–reduced allogeneic RBCs, compared with the recipients of autologous or WBC-reduced allogeneic RBCs. Moreover, as already discussed, if allogeneic WBCs are assumed to mediate the deleterious TRIM effect(s), then the problem of limited statistical power is compounded by the problem of limited exposure to allogeneic WBCs in these RCTs in which buffy-coat–reduced RBCs were transfused to the treatment arm.2,21,22 The question that is pertinent to clinical transfusion practice in North America—that is, whether transfusion of standard allogeneic RBCs increases the risk of cancer recurrence—has not yet been addressed by RCTs.

Furthermore, the setting of colorectal cancer resection may be inappropriate for the detection of a deleterious TRIM effect on cancer recurrence. Allogeneic transfusion–associated immunomodulation can be expected to increase the recurrence rate of a resected malignancy if the growth of residual cancer cells is indeed controlled by immunologic mechanisms. The existence of a specific immune response to colorectal cancer cells has not been proven.18 Although it is possible to generate cytotoxic T cells in vitro that recognize antigens expressed by colorectal cancer cells, the relevance of these cytotoxic cells in tumor growth may be limited because of loss of the expression of HLA molecules and adhesion molecules on colorectal cancer cells.85,86 If further RCTs are to be conducted, it may be preferable to concentrate on tumors known or presumed to be virus-induced and/or tumors that occur with increased frequency in patients receiving high doses of immunosuppressive drugs for the prevention of organ allograft rejection (ie, skin cancers, lymphoma, cervical carcinoma, and Kaposi sarcoma, as well as vulvar, perineal, and renal tumors).57

The reports of the 3 RCTs also presented observational comparisons of the incidence of cancer recurrence between patients having or not having transfusion in each study. The results of these analyses were conflicting.21,22 In the study of Houbiers et al,22 blood transfusion was not associated with cancer recurrence, and there was a statistically significant association between blood transfusion and mortality from causes other than cancer. Busch et al21 observed a 10% increase in the risk of cancer recurrence with transfusion (versus without transfusion) that could be attributed to local tumor recurrences (as opposed to distant metastases); the authors ascribed this increase in risk to the factors that generate a need for transfusion, as opposed to a TRIM effect.38 Heiss et al observed an increased risk of cancer recurrence in recipients of allogeneic (or autologous and allogeneic) RBCs compared with patients not having transfusion or recipients of autologous RBCs. However, their study sample was small (n = 100), and these 4 groups of patients were unbalanced with regard to the levels of potential confounding factors.

Postoperative infection

The 7 RCTs that compared the incidence of postoperative infection between recipients of buffy-coat–reduced or standard allogeneic RBCs or whole blood and recipients of autologous or WBC-reduced, buffy-coat–reduced allogeneic RBCs or whole blood were medically and statistically heterogeneous (Table 3). When all 7 studies were considered together, the probability that the disagreements among the findings of these studies might have arisen by chance was smaller than 1 per 10 000 (P < .0001 for the Q test statistic) (Figure 4). Two studies reported a significant (P < .05) TRIM effect, 2 studies reported a marginally significant (P < .10) effect, and 3 studies did not detect an effect. More important, the variation in reported results ranged from a 7.3-fold increase in the risk of infection (Figure 4) to no TRIM effect.2,22,23 Table 4 stratifies the 7 RCTs according to various study characteristics and then combines the results of RCTs that share each particular design attribute. The purpose of this meta-analysis is to examine whether specific differences in the design of the studies may have been responsible for the disagreements among the results. A study design attribute is considered to account for the disagreements among the studies if the integrated results from RCTs that have the same level of that attribute (eg, multicenter study organization) are statistically homogeneous and do not show an association between allogeneic transfusion and postoperative infection; and if the combined findings from RCTs that have other levels of the attribute (eg, single-center study organization) are also statistically homogeneous and demonstrate a statistically significant relationship between allogeneic transfusion and postoperative infection.

Five study characteristics could account individually for the disagreements among the 7 RCTs.3,7,22,23 These are: (1) study organization (ie, single center versus multicenter), (2) homogeneity of enrolled patients having elective abdominal surgery, (3) postoperative infection rate recorded in the entire study population, (4) proportion of patients having transfusion, and (5) study authorship. The RBC product transfused to the treatment or control arm could not account for the disagreements among the studies (Table 4). Single-center RCTs,3,7 studies enrolling a mixed patient population with any disease of the colorectum or gastrointestinal tract,3,5,7 studies recording an overall infection rate of 20% or less,3,7 trials in which less than 50% or more than 67% of the enrolled patients received perioperative transfusion,4,7 and RCTs reported by the team of Jensen et al were all homogeneous when considered together (P > .05 for the Q test statistic) and indicated a beneficial effect of the transfusion of autologous or WBC-reduced allogeneic RBCs. Multicenter RCTs,2,23 studies enrolling a homogeneous patient population having colorectal cancer resection,4,22,23 studies recording an overall infection rate of greater than 20%,22,23 trials in which 50% to 67% of the patients had transfusion,3,22,23 and RCTs reported by teams of investigators other than Jensen et al2,6,7,22,23 were also homogeneous when considered together (P > .05 for the Q test statistic) and indicated no difference in the incidence of postoperative infection between the treatment and control arms (Table 4).

Only the 2 multicenter RCTs22,23 reported an overall infection rate exceeding 20%; the overall infection rate was 20% or less in the 5 single-center studies. The results of the 2 multicenter RCTs of Busch et al23 and Houbiers et al22 may thus differ from the findings of the other studies because of the lower statistical power of multicenter (compared with single-center) studies to detect a treatment effect.99,101 In addition, Busch et al23 and Houbiers et al22 may have used more lenient criteria for the diagnosis of postoperative infection (or may have included more patients with risk factors for postoperative infection) compared with the other studies. Because the reports of all 7 RCTs presented only very limited information about the prevalence of risk factors for postoperative infection in the treatment and control arms of patients, the latter possibility cannot be investigated without access to the raw data.
variables that can influence the postoperative infection rate, but are
refers to differences among the participating medical centers in
Figure shows the OR of postoperative infection in recipients of buffy-coat–reduced 4–6,22,23
or standard 7 allogeneic RBCs or whole blood 3 as compared with recipients of
operative allogeneic blood transfusion with postoperative bacterial infection in
Figure 4. Randomized controlled trials investigating the association of periop-
Figure 4. Randomized controlled trials investigating the association of periop-
The “center effect” can be removed from the results of a multi-
intracohort variations among the participating medical centers in
cause of postoperative wound infections, because the incidence of
WBC-reduced blood transfusion effect and observed an extremely low postoperative
Jensen et al 3,5 reported an implausibly large allogeneic blood
Jensen et al3–5 reported an implausibly large allogeneic blood
not captured (or standardized) by the common research protocol.
The “center effect” can be removed from the results of a multi-
centered RCT if a statistical analysis stratified by participating
hospital is conducted, but such an analysis was not presented in the
studies of Jensen et al 3,5 and Tartter et al, 7 the OR is shown not as reported by the
autologous or WBC-reduced allogeneic RBCs or whole blood (Table 3). For the
*Used in the analysis of postoperative bacterial infection after patient exclusions and withdrawals.
†In both the treatment and control arms combined.
‡Investigators making diagnoses of postoperative infection were blinded to the treatment allocations.

With regard to the former possibility, multicenter RCTs often have
reduced statistical power compared with single-center studies
because of the influence of the “center effect” and their greater
susceptibility to the effects of random error. 89–92 The “center effect”
refers to differences among the participating medical centers in
variables that can influence the postoperative infection rate, but are
infection rate among the patients who received WBC-reduced
transfusion†

<table>
<thead>
<tr>
<th>Study</th>
<th>Total sample size*</th>
<th>Design</th>
<th>RBC product given to the treatment arm</th>
<th>RBC product given to the control arm</th>
<th>Overall percentage having transfusion†</th>
<th>Overall percentage developing postoperative infection†</th>
</tr>
</thead>
<tbody>
<tr>
<td>Jensen et al3</td>
<td>197</td>
<td>?Single blind</td>
<td>Allogeneic whole blood</td>
<td>WBC-reduced allogeneic whole blood (filtered after storage)</td>
<td>52.8</td>
<td>8.1</td>
</tr>
<tr>
<td>Heiss et al4</td>
<td>120</td>
<td>Unblinded</td>
<td>BC-reduced allogeneic RBCs</td>
<td>Autologous RBCs</td>
<td>75.0</td>
<td>20.0</td>
</tr>
<tr>
<td>Busch et al23</td>
<td>470</td>
<td>Unblinded</td>
<td>BC-reduced allogeneic RBCs</td>
<td>Autologous whole blood</td>
<td>59.0</td>
<td>26.0</td>
</tr>
<tr>
<td>Houbiers et al22</td>
<td>697</td>
<td>Single blind‡</td>
<td>BC-reduced allogeneic RBCs</td>
<td>WBC-reduced, BC-reduced allogeneic RBCs (filtered before storage)</td>
<td>64.0</td>
<td>33.4</td>
</tr>
<tr>
<td>Jensen et al5</td>
<td>589</td>
<td>Single blind‡</td>
<td>BC-reduced allogeneic RBCs</td>
<td>WBC-reduced, BC-reduced allogeneic RBCs (filtered after storage)</td>
<td>44.1</td>
<td>17.5</td>
</tr>
<tr>
<td>van de Watering et al6</td>
<td>909</td>
<td>Single blind‡</td>
<td>BC-reduced allogeneic RBCs</td>
<td>WBC-reduced, BC-reduced allogeneic RBCs (filtered before [n = 302] or after [n = 302] storage)</td>
<td>94.8</td>
<td>19.3</td>
</tr>
<tr>
<td>Tartter et al7</td>
<td>221</td>
<td>Unblinded</td>
<td>Allogeneic RBCs</td>
<td>WBC-reduced allogeneic RBCs</td>
<td>26.7</td>
<td>16.7</td>
</tr>
</tbody>
</table>

RBC indicates red blood cell; WBC, white blood cell; BC, buffy coat.
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Vamvakas and McAlister excluded the RCTs of Jensen et al. from their meta-analyses because of concerns about bias and residual, uncontrolled confounding factors in those studies. Jensen et al. published 2 of the 3 studies that enrolled a mixed population of patients having abdominal surgery for either benign or malignant disease. In addition, 2 of the 3 studies that enrolled a homogeneous population of patients having colorectal cancer resection were multicenter. Furthermore, of the 3 studies in which 50% to 67% of the enrolled patients had transfusion involved multiple centers. The influence of these study attributes (ie, homogeneity of the enrolled patient population and overall percentage of transfusion) on the reported results is probably due to the distribution of the 2 multicenter RCTs and the 2 trials by Jensen et al. among the categories formed when these study attributes are considered (Table 4).

To explain the disagreements among the 7 RCTs (Figure 4), Blajchman and Vamvakas and Blajchman proposed a meta-analysis of the 7 studies using individual patient data. The Table 4. Possible sources of variation in the findings of the available randomized controlled trials investigating the association between perioperative allogeneic blood transfusion and postoperative bacterial infection: integration of the findings of studies sharing a particular design characteristic by a meta-analysis

<table>
<thead>
<tr>
<th>Study attribute</th>
<th>Categories of study attribute</th>
<th>Number of studies</th>
<th>Q test statistic: P value</th>
<th>Summary odds ratio*</th>
<th>Point estimate</th>
<th>95% confidence interval</th>
</tr>
</thead>
<tbody>
<tr>
<td>Study organization</td>
<td>Single center</td>
<td>3-7</td>
<td>5</td>
<td>&gt; .05</td>
<td>2.29†</td>
<td>1.35-3.87†</td>
</tr>
<tr>
<td></td>
<td>Multicenter</td>
<td>22,23</td>
<td>2</td>
<td>.95</td>
<td>.88</td>
<td>0.69-1.12</td>
</tr>
<tr>
<td>Surgical setting</td>
<td>Open heart surgery</td>
<td>6</td>
<td>1</td>
<td>N/A†</td>
<td>1.42‡</td>
<td>1.00-1.99‡</td>
</tr>
<tr>
<td></td>
<td>Elective abdominal surgery</td>
<td>3,5,22,23</td>
<td>6</td>
<td>&lt; .001</td>
<td>N/A</td>
<td>N/A</td>
</tr>
<tr>
<td>Homogeneity of patients having elective abdominal surgery: diagnosis</td>
<td>Colorectal cancer</td>
<td>4,22,23</td>
<td>3</td>
<td>&gt; .10</td>
<td>0.96</td>
<td>0.67-1.37</td>
</tr>
<tr>
<td></td>
<td>Any disease of the colorectum or gastrointestinal tract</td>
<td>3,5</td>
<td>3</td>
<td>&gt; .25</td>
<td>2.95†</td>
<td>1.73-5.05†</td>
</tr>
<tr>
<td>RBC product transfused to the treatment arm</td>
<td>Allogeneic whole blood</td>
<td>3</td>
<td>1</td>
<td>N/A†</td>
<td>7.32¶</td>
<td>1.62-33.12¶</td>
</tr>
<tr>
<td></td>
<td>Allogeneic RBCs</td>
<td>7</td>
<td>1</td>
<td>N/A†</td>
<td>1.85¶</td>
<td>0.89-3.85¶</td>
</tr>
<tr>
<td></td>
<td>Buffy-coat-reduced RBCs</td>
<td>4,22,23</td>
<td>5</td>
<td>&lt; .001</td>
<td>N/A</td>
<td>N/A</td>
</tr>
<tr>
<td>RBC product transfused to the control arm</td>
<td>Autologous RBCs or whole blood</td>
<td>23</td>
<td>2</td>
<td>&gt; .05</td>
<td>1.35</td>
<td>0.45-4.08</td>
</tr>
<tr>
<td></td>
<td>WBC-reduced allogeneic RBCs</td>
<td>5,7,22,23</td>
<td>5</td>
<td>&lt; .001</td>
<td>N/A</td>
<td>N/A</td>
</tr>
<tr>
<td></td>
<td>WBC-reduced allogeneic RBCs</td>
<td>5,7,22,23</td>
<td>4</td>
<td>&lt; .001</td>
<td>N/A</td>
<td>N/A</td>
</tr>
<tr>
<td></td>
<td>Fresh-filtered, WBC-reduced allogeneic RBCs</td>
<td>22</td>
<td>2</td>
<td>&lt; .05</td>
<td>N/A</td>
<td>N/A</td>
</tr>
<tr>
<td></td>
<td>WBC-reduced allogeneic RBCs</td>
<td>22</td>
<td>3</td>
<td>&lt; .01</td>
<td>N/A</td>
<td>N/A</td>
</tr>
<tr>
<td>Overall postoperative infection rate#</td>
<td>Exceeding 20%</td>
<td>22,23</td>
<td>2</td>
<td>.95</td>
<td>0.88</td>
<td>0.69-1.12</td>
</tr>
<tr>
<td></td>
<td>20% or lower</td>
<td>3,5</td>
<td>5</td>
<td>&gt; .05</td>
<td>2.29†</td>
<td>1.35-3.87†</td>
</tr>
<tr>
<td>Overall percentage of patients having transfusion#</td>
<td>Less than 50%</td>
<td>5,7</td>
<td>2</td>
<td>&gt; .10</td>
<td>2.73†</td>
<td>1.48-5.04†</td>
</tr>
<tr>
<td></td>
<td>50% to 67%</td>
<td>3,22,23</td>
<td>3</td>
<td>&gt; .10</td>
<td>0.96</td>
<td>0.62-1.49</td>
</tr>
<tr>
<td></td>
<td>More than 67%</td>
<td>4,6</td>
<td>2</td>
<td>&gt; .25</td>
<td>1.53†</td>
<td>1.00-2.33†</td>
</tr>
<tr>
<td>Reported by Jensen et al. or by other teams of investigators</td>
<td>Jensen et al.</td>
<td>3,5</td>
<td>2</td>
<td>.50</td>
<td>3.68</td>
<td>2.11-6.43</td>
</tr>
<tr>
<td></td>
<td>Other teams</td>
<td>4,6,7,22,23</td>
<td>5</td>
<td>&gt; .05</td>
<td>1.17</td>
<td>0.84-1.62</td>
</tr>
</tbody>
</table>

RBC(s) indicates red blood cell(s); WBC, white blood cell; N/A, not applicable.
*Calculated by the method of DerSimonian and Laird.
†Statistically significant (P < .05) summary odds ratio.
‡Not applicable because only one study was available for meta-analysis.
§Odds ratio of postoperative infection in patients receiving buffy-coat-reduced RBCs, as compared with recipients of WBC-reduced, buffy-coat-reduced RBCs filtered before or after storage.
¶Odds ratio shown not as reported by the authors, but as recalculated according to an intention-to-treat analysis.
#In the entire study population regardless of randomization arm.
proposed project would require recoding of the raw data collected prospectively by the authors of the 7 RCTs with a common patient data form, as well as the collection of additional data through a retrospective review of the medical records of the patients who were enrolled in the 7 studies. The additional information needed to explain the disagreements among the studies relates primarily to the prevalence of risk factors for postoperative infection, as well as the criteria used to reach each diagnosis of postoperative infection in the 7 RCTs. If information on these variables were obtained, it would highlight any differences in severity of illness or application of the diagnostic criteria for postoperative infection between the treatment and control arms of the studies, thereby allowing meta-analysts to assess the possible effects of bias and confounding in the 7 RCTs. In addition, the variation in the postoperative infection rate among the studies (Table 1) could be removed, or attributed to a differential distribution of risk factors for postoperative infection among the RCTs; and the data from the 2 multicenter RCTs could be stratified by participating hospital to remove the influence of the “center effect.” If these postulated sources of variation accounted for the disagreements among the studies, a summary estimate of the TRIM effect (or of the benefit from the use of autologous and/or WBC-reduced RBCs) across the 7 RCTs could be calculated by the techniques of meta-analysis.

Because the results of an IPD meta-analysis of all published RCTs are not available, the results of the study of van de Watering et al may be the most reliable existing data regarding the potential benefit from the use of WBC-reduced blood components for the prevention of postoperative infection. The study of van de Watering et al was well designed and analyzed and enrolled a large number of patients, almost all of whom received transfusion. Furthermore, observers who were unaware of patient assignment made the diagnoses of postoperative infection on the basis of the detailed criteria of the Centers for Disease Control and Prevention. All of these conditions together were not met by any of the other RCTs shown in Table 3.

Van de Watering et al randomized 914 patients scheduled to have CABG surgery, cardiac valve surgery, or a combination of the 2 to one of 3 treatment arms: buffy-coat–reduced allogeneic RBCs; WBC-reduced, buffy-coat–reduced allogeneic RBCs filtered before storage; and WBC-reduced, buffy-coat–reduced allogeneic RBCs filtered after storage. The intention-to-treat analysis of the incidence of postoperative infection among these 3 arms and found no significant difference (P = .13); the incidence of postoperative infection was 23%, 16.9%, and 17.9%, respectively, in each arm. However, when the prestorage and poststorage WBC-reduced arms were combined, recipients of buffy-coat–reduced RBCs had a higher incidence of postoperative infection than patients receiving WBC-reduced, buffy-coat–reduced RBCs, and this difference was marginally significant (P = .06; Figure 4).

**Postoperative mortality**

In addition to a possible association between allogeneic transfusion and postoperative infection (Figure 4), van de Watering et al detected an unexpected association between WBC-containing allogeneic blood transfusion and postoperative mortality from causes other than postoperative infection. Twenty-four of 306 patients (7.8%) having transfusion with buffy-coat–reduced RBCs died, compared with 11 of 305 patients (3.6%) receiving buffy-coat–reduced RBCs that were WBC reduced before storage and 10 of 303 patients (3.3%) receiving buffy-coat–reduced RBCs that were WBC reduced after storage (P = .015). This overall difference in 60-day mortality was due to a highly significant (P = .001) difference among the 3 randomization arms in the proportion of patients who died of noncardiac causes, that is, multiorgan failure or dehiscence of their aortic bypass anastomosis (3.3% of patients receiving buffy-coat–reduced RBCs, as compared with 0.3% of patients receiving buffy-coat–reduced RBCs that were WBC reduced either before or after storage). The proportions of patients who died of cardiac causes (ie, myocardial infarction, heart failure, or arrhythmia) did not differ among the 3 arms (P = .53).

Furthermore, in an analysis of the effect of the volume of blood infused among the 866 patients who had transfusion, no difference in mortality between recipients of buffy-coat–reduced or WBC-reduced RBCs was observed in the subgroup of subjects receiving 1 to 3 U of blood (P = .82). In this subgroup of 355 patients, mortality was 1.6% for the recipients of buffy-coat–reduced units, as compared with 1.8% for the recipients of WBC-reduced units filtered after storage and 0.9% for the recipients of WBC-reduced units filtered before storage. In contrast, in the subgroup of 511 patients who received 4 U or more of RBCs, there was a statistically significant (P = .014) association between mortality and WBC-containing allogeneic blood transfusion across the 3 arms. When the recipients of WBC-reduced RBCs filtered before or after storage were combined, the postoperative mortality rate of recipients of buffy-coat–reduced RBCs was 12.5%, as compared with 5.1% for the recipients of WBC-reduced, buffy-coat–reduced RBCs (P = .005).

Van de Watering et al pointed out that their study had not been designed to investigate postoperative mortality as an outcome variable. For this reason, prognostic factors for mortality had not been measured during the trial, and randomization might (or might not) have distributed such prognostic variables equally among the 3 arms. Accordingly, using multivariate regression analysis, van de Watering et al adjusted the observed difference across the randomization arms for the effects of measured prognostic factors for mortality. All variables that demonstrated a univariate association (P < .05) with postoperative mortality were entered into this multivariate regression model. Randomization to the buffy-coat–reduced RBC arm was associated with significantly increased mortality (P = .012) after adjustment for the effects of type of operation, age, history of myocardial infarction, previous open heart surgery, preoperative platelet count, and gender. When the number of RBC units transfused was entered into the regression model, it was the most significant (P < .001) predictor of postoperative mortality, while randomization to the buffy-coat–reduced RBC arm still exercised a significant (P = .009) effect on mortality.

Van de Watering et al concluded that when 4 U or more of allogeneic buffy-coat–reduced RBCs are transfused in cardiac surgery, WBC reduction of the transfused units results in significantly decreased mortality. However, recognizing that this finding was novel and unexpected, and that a recommendation for using only WBC-reduced blood components in cardiac surgery would significantly increase the cost of transfusion, these investigators urged that further research be conducted to replicate their finding before such a recommendation was made.

Van de Watering et al postulated that their study was more likely than previous RCTs to detect differences among the randomization arms because the transfusion dose administered to their cardiac surgery patients was substantially higher than that given to patients having colorectal surgery in the previously published trials. Only 48 (5.2%) of these cardiac surgery patients received no RBC transfusion; 777 (85.0%) received 1 to 10 U of RBCs each (mean, 5.5 U; median, 4.0 U) and 89 (9.7%) received more than 10 U each (mean, 16.9 U; median, 14.0 U). To further
investigate a possible relation between WBC-containing allogeneic blood transfusion and an increased risk of death or multiorgan failure in patients receiving high transfusion doses perioperatively, investigators from the Leiden University Medical Center initiated an RCT enrolling subjects admitted for resection of aortic aneurysm or gastrointestinal malignancy. The purpose of the new study is to determine whether the previously detected associations between WBC-containing allogeneic blood transfusion and increased risk of death or multiorgan failure represent transfusion-associated complications (TAC) or transfusion-induced complications (TICS); thus, the acronym TACTICS is used to refer to this new study.99 To optimize the design of this large, double-blind, multicenter trial, Brand et al100 conducted a pilot study enrolling 73 patients with aortic aneurysm and 84 patients with abdominal malignancy. No statistically significant differences were observed in the pilot study. Ten percent of the recipients of buffy-coat–reduced RBCs died, as compared with 7.8% of the recipients of WBC-reduced RBCs; approximately 20% of the patients in either randomization arm developed multiorgan failure.100

Information on 30-day mortality was available for 1806 patients enrolled in the 3 RCTs conducted at the Leiden University Medical Center (ie, the RCTs of Houbiers et al22 and van de Watering et al6 and the pilot component of TACTICS), and an interim analysis based on the accumulated experience from these patients has been presented.99,100 In the univariate analysis, an increased mortality risk in recipients of buffy-coat–reduced (compared with WBC-reduced) RBCs was observed only among recipients of 4 U or more of RBCs. In a multivariate regression analysis, randomization to the WBC-reduced (as compared with the buffy-coat–reduced) arm was associated with a statistically significant decrease in postoperative mortality after adjustment for the effects of age, gender, and transfusion dose. However, because prognostic factors for mortality had not been considered in the design of the 3 combined studies, there may exist other prognostic variables that could be unequally distributed between the buffy-coat–reduced and WBC-reduced arms and that could account for the observed difference in mortality between the arms.

The association between WBC-containing allogeneic blood transfusion and increased mortality92 may have limited applicability because data obtained from patients undergoing cardiac surgery may not be generalizable to other clinical settings. The extracorporeal circuit used in cardiac surgery induces a diffuse inflammatory response that may predispose to postoperative infection or other surgical complications.101 It is possible that WBC reduction of transfused blood components may be capable of suppressing the diffuse inflammatory response to the extracorporeal circuit and may thus indirectly reduce the incidence of postoperative infection or the incidence of other postoperative complications.101 If this were the mechanism of the beneficial effect observed by van de Watering et al,6 the same benefit from the use of WBC-reduced blood components would probably not be expected to occur in other surgical settings.

The mechanism of the reported association6 between WBC-containing allogeneic blood transfusion and increased postoperative mortality from causes other than postoperative infection is unclear. If such a relationship truly exists, its mechanism may (or may not) be related to transfusion-induced immunomodulation. It is possible that, among recipients of multiple transfusions, allogeneic transfusion may exercise other deleterious effects that have not yet been recognized or defined, probably because of the confounding effect of the association between transfusion and illness severity (which becomes extremely strong in patients receiving multiple transfusions). Associations between allogeneic transfusion and various adverse clinical outcomes (eg, prolonged mechanical ventilation,102 impaired wound healing,103 or multiorgan failure104-107) have been reported from retrospective cohort studies, but it remains uncertain whether these associations reflect true relationships because it has been impossible to separate the effects of the allogeneic transfusions from the effects of confounding factors.

Hébert et al52 reported that a restrictive strategy of (non–WBC-reduced) RBC transfusion may be superior to a liberal transfusion strategy in critically ill patients with normovolemia. Thirty-day mortality was 18.7% in patients having transfusion when their hemoglobin concentration fell to 7 g/dL, as compared with 23.3% in subjects having transfusion when their hemoglobin concentration fell to 10 g/dL (P = .11). In subgroup analyses, patients who were less acutely ill (with an Acute Physiology and Chronic Health Evaluation II score of ≤ 20) and patients who were younger than 55 years experienced a statistically significant survival benefit by having been allocated to the restrictive-strategy arm of the study. Among the less acutely ill patients, 30-day mortality was 8.7% in the restrictive-strategy arm and 16.1% in the liberal-strategy arm (P = .03). Among the patients who were younger than 55 years, 30-day mortality was 5.7% with the restrictive strategy and 13.0% with the liberal strategy (P = .02).

As with the finding of a reduced 60-day mortality rate among the recipients of WBC-reduced RBCs reported by van de Watering et al,6 the finding of a trend (P = .11) toward reduced 30-day mortality in the restrictive-strategy arm of the RCT of Hébert et al52 was unexpected. There were no differences between the arms in the rates of cardiac events, infectious complications, multiorgan failure of more than 3 organs, or multiorgan failure in the 48 hours before death. The harm caused by the liberal use of transfusions in critically ill patients was thus difficult to explain.108 It is possible that the overall trend observed by Hébert et al52 was due to chance because a statistically significant difference between the restrictive- and liberal-strategy arms could be demonstrated only in subgroup analyses. Alternatively, it is possible that allogeneic transfusion may exercise other deleterious effects that have not yet been recognized or defined. More research is needed to elucidate the possible biologic mechanism(s) that may underlie such other adverse effects of allogeneic transfusion.27

Conclusions

A causal relationship between allogeneic blood transfusion and cancer recurrence or postoperative infection has not been proven by double-blind RCTs in accordance with the tenets of evidence-based medicine.109 The available RCTs2,21,22 provide no indication that perioperative allogeneic blood transfusion causes an increase in cancer-related complications, whether cancer recurrence or death due to cancer recurrence is considered the outcome variable. The OR of an adverse outcome in the allogeneic blood transfusion (compared with the control) arm is approximately 1. In the case of postoperative infection, the results of the available RCTs3,7,22,23 are highly contradictory, and they cannot be combined by the techniques of meta-analysis to generate an estimate of the TRIM effect based on all available data.

Observational studies comparing recipients of similar volumes of autologous or allogeneic blood indicate an increased risk of postoperative infection in association with allogeneic transfusion (Figure 2). However, it is impossible to determine from the published results whether some portion of this excess risk seen in association with allogeneic transfusion would persist if the effects of selection bias, observation bias, and confounding factors were to
be removed completely. Observational studies comparing patients with and without transfusion, and using multivariate analysis to adjust for the effects of confounding factors, almost uniformly support the hypothesis of an increased risk of postoperative infection in the transfusion group. However, in most of these studies, the reported allogeneic transfusion effects were not adjusted for the effects of illness severity and risk factors for postoperative infection at specific sites.

Reported RCTs investigating the association of perioperative allogeneic blood transfusion with postoperative infection were either unblinded or single-blind (Table 1), and no double-blind RCTs have been reported. Because the diagnosis of nosocomial infections is subjective, observation bias is an important concern, and it may represent an alternative explanation for the findings of some studies and for disagreements among the reported RCTs. Furthermore, the reports of the available RCTs did not present sufficient information about illness severity or the distribution of risk factors for postoperative infection in the treatment versus control arms. In the absence of this information, the possible contributions of selection bias, observation bias, and confounding factors to the reported results cannot be evaluated, and therefore cannot be excluded.

A possible explanation for the disagreements among the 7 RCTs may be that the adverse TRIM effect is small: for example, a 5% to 10% increase in the risk of postoperative infection in association with allogeneic blood transfusion. Such a small effect would not be detected consistently, and its detection would be highly dependent on the size and the design of each study. To detect a 10% difference in the risk of postoperative infection between treatment and control arms of equal size, 3000 patients would have to be randomized to each arm if the overall infection rate in the entire study population were 35%. A study population of 20,000 patients would be required if the overall infection rate were approximately 20% and if almost half of the enrolled patients did not receive transfusion. Therefore, the available data from the 3203 patients enrolled in the 7 published RCTs are grossly insufficient for the investigation of an adverse TRIM effect of such small magnitude. In fact, it is unlikely that a definitive RCT capable of detecting such a small effect will ever be conducted. To reach a conclusion about the existence and magnitude of any increase in the risk of postoperative infection that is caused by allogeneic transfusion, decision makers will probably have to rely on cumulative meta-analyses of the data from all available RCTs, and the excessive variation in the findings of the published reports will have to be accounted for to conduct such meta-analyses.

As shown in Table 4, it is impossible to reconcile the findings of all the available RCTs by relying only on the published information. The crux of the disagreements involves the radically different results reported by the team of Jensen et al and the 2 multicenter RCTs. To permit a meta-analytic synthesis of the published findings, either the studies of Jensen et al or the 2 multicenter RCTs have to be excluded from the analysis. In the absence of objective evidence of bias or confounding, any exclusions are arbitrary, and they bias the conclusions of the meta-analysis either in the direction of a null result or in the direction of a large (and likely spurious) deleterious TRIM effect (Table 4). A reanalysis of the raw data on the individual patients enrolled in the 7 RCTs and collection of additional information on these subjects are needed to reconcile the published results and to reach a conclusion about the existence of a deleterious TRIM effect based on all available data from clinical experimental studies.

The question currently debated at the national level is the issue of universal WBC reduction. If TRIM really causes deleterious effects in recipients of blood products, these likely affect many (or potentially all) patients having transfusion. Therefore, it would be logical to implement measures to prevent these immunomodulatory effects of allogeneic blood transfusion. Because most evidence from studies in experimental animals implicates allogeneic WBCs in the production of the deleterious TRIM effects, and because universal WBC reduction of all transfused cellular blood components has been implemented in some western European countries and in Canada, policy makers in the United States and elsewhere are currently considering the question of implementing universal WBC reduction for the purpose of preventing the various possible deleterious TRIM effects.

As shown in Table 4, when the 5 RCTs that transfused WBC-reduced allogeneic RBCs or whole blood to the control arm were considered together, the results were heterogeneous (P < .001 for the Q test statistic) and could not be combined. Therefore, the existing evidence regarding the efficacy of WBC reduction in preventing the deleterious TRIM effects is conflicting. If standards similar to those used by the FDA’s new-drug evaluation process were applied to transfusion safety, the only appropriate recommendation at this time would be that further double-blind RCTs be conducted to establish the existence of any adverse TRIM effects mediated by transfused allogeneic WBCs, as well as the efficacy of WBC reduction in preventing these effects.

Decisions made in the 1990s, however, indicate a willingness on the part of policy makers to use a substantially different threshold for judging the efficacy of measures that affect transfusion safety, compared with the criteria used for the evaluation of other therapeutic interventions. Because of these precedents, the question now is whether universal WBC reduction should be implemented in the United States and elsewhere at approximately the same time that this practice is introduced in western Europe and Canada.

Should universal WBC reduction be introduced now, or should we wait for definitive, large, double-blind RCTs to establish both the existence of deleterious TRIM effects and the efficacy of WBC reduction in abrogating these effects? Table 5 looks at this question from a public policy (as opposed to an evidence-based medicine) perspective and summarizes the arguments for and against a policy to implement universal WBC reduction for the prevention of the adverse TRIM effects, based on the current state of knowledge. All “pro” arguments are logical, and they are countered by equally reasonable “con” arguments. Therefore, any decision at this time to implement (or not to implement) universal WBC reduction in the United States for the purpose of preventing the purported deleterious TRIM effects will have to be arbitrary. One can argue that a policy decision should be made on the basis of the existing evidence because better evidence is unlikely to be forthcoming. Alternatively, one can argue that no policy decision should be made until it is possible to make such a decision based on sound scientific evidence.

Regardless of the content of the public policy debate, should universal WBC reduction be implemented, the question as to whether WBC-containing allogeneic blood transfusion causes deleterious TRIM effects will remain. Therefore, should a decision to introduce universal WBC reduction be made in the near future, research into the purported deleterious TRIM effects should continue. Double-blind RCTs should be conducted in countries that do not implement universal WBC reduction; observational studies comparing the incidence of postoperative infection and/or cancer
Table 5. Arguments for (pro) and against (con) a policy to implement universal white cell reduction of all transfused cellular blood components for the purpose of preventing postoperative infection caused by the purported deleterious immunomodulatory effects of allogeneic blood transfusion

<table>
<thead>
<tr>
<th>Pro26,27</th>
<th>Con27</th>
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<tr>
<td>1. An increase in the incidence of postoperative infection attributable to an immunomodulatory effect of allogeneic WBCs is supported by at least one single-blind and correctly designed and analyzed RCT conducted by van de Watering et al.5</td>
<td>1a. The RCT of van de Watering et al5 did not detect a TRIM effect in the intention-to-treat analysis; only when the 2 control (WBC-reduced) arms were combined was a marginally significant (P = .06) difference demonstrated between the recipients of buffy-coat–reduced and WBC-reduced RBCs.</td>
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<tr>
<td>2. Four2,5,7 of 5 RCTs3,5,7,22 investigating the relationship between WBC-containing (versus non–WBC-containing) allogeneic blood transfusion and postoperative infection show at least a trend toward an increase in the incidence of postoperative bacterial infection in association with the transfusion of allogeneic WBCs.</td>
<td>1b. The RCT of van de Watering et al5 enrolled patients having open heart surgery, and an allogeneic blood transfusion effect detected in the setting of cardiac surgery may not apply to other surgical settings.</td>
</tr>
<tr>
<td>3. It is unlikely that further RCTs examining the relationship between allogeneic blood transfusion and postoperative infection will be reported in the foreseeable future. Even if more studies are conducted in the future, it is unlikely that such future RCTs will be large enough to be definitive.</td>
<td>2. The 2 RCTs by Jensen et al5,8 reported an implausibly large allogeneic blood transfusion effect, and, if these 2 studies3,5 were excluded from the analysis, the OR of postoperative infection in the allogeneic transfusion (compared with the control) arm across the remaining RCTs6,7,22 would no longer be statistically significant and would be sufficiently modest to be attributed to chance.6</td>
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<td>4. In the future, research into the purported deleterious immunomodulatory effects of allogeneic blood transfusion can continue, by means of observational studies comparing the incidence of postoperative infection between patients having transfusion before or after the implementation of universal WBC reduction.</td>
<td>3. If a policy decision is made to implement universal WBC reduction based on the evidence that is currently available, it will be impossible to rescind this policy in the future if further studies establish that universal WBC reduction does not decrease the incidence of postoperative infection.</td>
</tr>
<tr>
<td>5. The receipt of WBC-containing allogeneic blood transfusion is associated with increased postoperative mortality in the RCT of van de Watering et al.5</td>
<td>4. Such future observational studies are bound to have at least some of the flaws of their predecessors,19 making it impossible to attribute any observed difference between patients having transfusion before or after the implementation of universal WBC reduction to the receipt of WBC-containing allogeneic blood transfusion.</td>
</tr>
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</table>

WBC indicates white blood cell; RCT, randomized controlled trial; TRIM, allogeneic blood transfusion–associated immunomodulation; RBC, red blood cell; OR, odds ratio. *The OR of postoperative infection in recipients of buffy-coat–reduced26,27 or standard RBCs, as compared with recipients of WBC-reduced RBCs, was 1.20 (95% confidence interval for the OR, 0.78–1.85; P > .05). These 3 studies3,7,22 are sufficiently homogeneous to permit meta-analysis (P > .05 for the Q test statistic).

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Deleterious clinical effects of transfusion-associated immunomodulation: fact or fiction?

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