To the editor:

Postmenstrual gestational age should be used with care in studies of early human hematopoietic development

The hematopoietic system undergoes rapid changes during embryogenesis; therefore, studying this process requires accurate embryo staging. In the mouse, timed pregnancies can routinely be set and controlled; however, accurate staging of human pregnancies is more problematic, compounded by different ways that gestational age is assessed.

Although developmental biologists calculate gestational age based on the woman’s last ovulation (postovulatory gestational age), clinicians evaluate it from the first day of the last menstrual period (postmenstrual gestational age), using ultrasound measurement of crown-rump length. Recording the postmenstrual gestational age of the human embryo largely meets the needs of the medical care. However, it will exceed postovulatory gestational age, which is the true developmental age of the embryo, by 2 weeks in the normal case of a regular menstrual period (postmenstrual gestational age cannot be used to distinguish between CS 12–15, the 95% confidence interval were calculated for each of the groups (Figure 1). For embryos categorized as CS 12–15, the 95% confidence intervals of the postmenstrual gestational age overlapped between the 4 groups, indicating that postmenstrual gestational age cannot be used to distinguish between CS 12–15 human embryos. Also, there was no relationship between postmenstrual gestational age of individual CS 12–15 embryos and the CSs to which they were assigned (correlation coefficient $r = .251; P = .101$). However, CS 16 and 17 embryos could be distinguished based on their postmenstrual gestational age at a 95% confidence level, and there was a direct relationship between postmenstrual gestational age and the CSs to which they belonged ($r = .536; P < .001$).

These data demonstrate that postmenstrual gestational age cannot be used to accurately stage human embryos under CS 16. This is of particular importance for those working in the field of human developmental hematopoiesis because the first human hematopoietic stem cells emerge at CS 14.14

Figure 1. Relationship between postmenstrual gestational age and CSs. Human embryos obtained for the present study were split into groups depending on CSs to which they belonged. For each group, a mean postmenstrual gestational age in days ($) and a 95% confidence interval (error bars) were calculated and plotted against corresponding CSs. The number of human embryos obtained for each CS is indicated above the error bars.

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3051
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References

To the editor:

Utility in prognostic value added by molecular profiles for diffuse large B-cell lymphoma

We read with interest the report from Hong et al1 on the incremental prognostic value of gene expression signatures in diffuse large B-cell lymphoma (DLBCL). We were surprised by the conclusions that these “signatures are inferior to clinical factors and provide little added value in risk assessment” when considering a 2-gene score (TGS) and a 6-gene score described by us2-4 and others.5 The authors’ claim that “all studies assess predictive significance based on P value from multivariable Cox regression” was also surprising, because dedicated parts of prior studies specifically addressed added prognostic value. We reported robust risk reclassification by integration of molecular indices (TGS) with clinical factors within the international prognostic index (IPI) (TGS-IPI),2 and both our study and the study of Lenz et al5 demonstrated splitting of IPI strata using molecular signatures.

We favored risk reclassification as a measure of added value from molecular indices over more complex statistical tests for several reasons (pie charts in figure 4B of Alizadeh et al5). First, even statistically significant biomarkers may yield minimal improvement in the area under the curve.5-7 Therefore, many statisticians consider C-statistics unsuitable for assessing improvement in prediction, preferring the use of measures such as net reclassification improvement and integrated discrimination improvement (IDI).7-10 Second, C-statistics and IDI lack an intuitive interpretation for clinicians. However, most surprisingly, we came to the opposite conclusion as the authors in considering these indices in 3 cohorts of DLBCL patients (n = 561), including an important validation cohort that they did not consider. We provide these new data here (Figure 1A), having originally opted not to present them for space considerations.

Although Hong et al do not provide detailed methods to fully address this, several aspects of the analysis may have led to the discrepancy in their conclusions.

1. There is an error in the LIM domain only 2 (rhombotin-like 1) (LMO2) coefficient (0.032 instead of 0.32) for calculating the TGS.
2. The model parameters we reported weight gene expression values measured by quantitative real-time reverse-transcription polymerase chain reaction, with array data needing appropriate rescaling. The TGS values obtained by the authors are thus incorrect, because rescaling is equivalent to changing the relative weightings of the genes in the model, similar to how an equation considering “age” would differ depending on whether it was measured in days or years.
3. We used a custom gene-level chip definition to renormalize probe-level Affymetrix array data from Lenz et al,5 providing more accurate gene expression quantification (Figure 1B-C).
4. The publicly available normalized expression data used by the authors are already log2 transformed. Doing this again as suggested in the methods would result in data distortion (Figure 1D).
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